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Renewable Generation and Electricity Prices: Taking Stock and New Evidence for Germany and Austria

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Abstract

Economic theory predicts that high renewable electricity production reduces the price of electricity, also referred to as the ‘merit-order effect’. Although the merit-order effect is only one of several consequences of renewable production on the electricity system, it is crucial to determine its size for the economic evaluation of renewable energies. In this paper we present a comprehensive overview of relevant past research results on the price effect of renewables. Additionally, we conduct a new empirical analysis of the price effect of renewable production for the Austrian-German region, a market that clearly qualifies for a merit-order effect analysis given its characteristics. Based on the review and our own analysis, we show that the merit-order effect varies depending on the region and the assessment method chosen. We also find that the size of this effect is less dispersed throughout different markets than previously suggested by the literature.

Keywords: Wind, solar, merit-order, supply

JEL classification : Q41, Q42, Q48

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1. Introduction

Renewable energy sources are a key element in strategies and scenarios to tackle climate change (see e.g. IPCC, 2011; IEA, 2012), and they have also been vindicated as a way to reduce energy dependence or to promote new economic sectors and activities in specific countries. Yet it is usually more expensive to generate electricity with renewables than it is with conventional (mostly fossil-fuel) technologies. Indeed, countries that have succeeded in increasing their renewable capacities significantly have done so by implementing intense support policies, mostly feed-in-tariffs (FIT) or renewable allowance systems1, which incur in significant costs. In many cases the very success of the renewable promotion schemes has led to increasing costs, such as with FIT, thus raising doubts on their future economic viability. This has been recently the case in the two most ambitious European experiences in the field so far, with a moratorium on the whole Spanish renewable support system and severe amendments to counter cost increases in the German FIT system2.

Identifying and quantifying the costs and benefits of increased renewable capacity is therefore a key element of economic research on renewable energy. Indeed, the impacts of additional renewable electricity production in energy markets must be fully understood to evaluate the efficiency (i.e. desirability) of renewable support policies. One very important question in this regard is whether renewable production affects the electricity price and, if so, how it does so. Theoretical considerations first stated by Jensen and Skytte (2002) suggested that renewable electricity production results in lower electricity prices, as subsequently elaborated by, among others, Sensfuss et al. (2008) or Nicolosi and Fürsch (2009). The phenomenon is commonly known as the ‘merit-order effect’, where price decreases occur because (additional) renewable electricity bids into the market at lower marginal costs.

From a political point of view these price reductions often seem appealing and they frequently serve as a justification for renewable support schemes. However, as already stated by Gelabert et al. (2011), the merit-order effect is mainly a transfer of wealth from producers to consumers, at

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1 Several European countries have managed to create a sizeable renewable capacity. See Schmalensee (2012) for a critical assessment of renewable support schemes in the EU and the US.
2 At the beginning of 2012 Spain introduced a moratorium of all renewable support after experiencing an unexpected increase in solar capacity, which imposed strong financial pressure on its FIT system. That spring, Germany also implemented unprecedented cuts in its support to photovoltaic technologies and, given the persistence of the problem, recently issued further restrictive proposals. For an updated description of the Spanish and German renewable support experiences see, respectively, Linares and Labandeira (2013) and Diekmann et al. (2012).
least if market equilibrium is formed at prices equal to marginal costs\(^3\), with gains and losses potentially shared unevenly among different types of generators (depending on their cost structures). Furthermore, a more intensive use of renewables requires an increase of conventional back-up capacity to cope with the volatility of renewable generation. This process may, in turn, affect long-run electricity prices in the opposite direction (Nicolosi and Fürsch, 2009; Jonsson et al., 2010), thus offsetting the initial (short-term) merit-order effect (see e.g. Green and Vasilakos, 2011). These factors, along with the potential conflict of interest between suppliers and consumers of electricity, may introduce controversy in the calculation of the benefits and costs of renewable energy.

Any program that aims to assess past regulation and ensure cost-effectiveness of future policies should therefore carry out a comprehensive and rigorous analysis of the price effects of renewables. There is a considerable literature on this issue now, so it is possible to take stock and draw some general conclusions. However, this is not an easy endeavor because the measurement of price effects depends on several diverging factors contemplated in the different studies (technological mix in the electricity system, market conditions etc.)\(^4\), as well as on the different methodologies employed to do so. Indeed, comparability constraints range from the use of different datasets or different systematic approaches (such as the use of simulations or empirical analysis of market data) to other small but possibly influential factors (e.g. the inclusion of additional explanatory factors such as the working load or the definition of the renewable variable: wind speed, wind generation, or a sum of wind and solar generation). Therefore, results must be interpreted conjointly and should reflect the particular differences among the studies whilst attempting to draw useful conclusions from existing research for a global analysis. We follow this approach in the first part of the paper, designed to both provide a detailed overview of the literature and reveal previously undisclosed patterns in prior research.

The paper also presents the results of an own, novel empirical investigation on the price effect of renewable production for the German-Austrian electricity market. We carried out this empirical exercise for several reasons, but we did so mostly as a way to complete and interpret the surveyed literature for a key European energy and economic area with limited empirical evidence

\(^3\) If firms have market power and can set prices above marginal costs, however, renewable-related price reductions may enhance total welfare. But even then, rents are shifted from producers to consumers in significant quantities.

\(^4\) Transmission capacity limits within the network of an electricity market may cause price discrepancies, such as those seen, for instance, in Texas (Zarnikau, 2011).
on this matter (particularly in empirical approaches). Moreover, the German-Austrian market features an important renewable capacity that is obviously related to the strong renewable support scheme that has been in place for many years. The fact that a market of this size, located in the core of Europe, may strongly influence other closely integrated energy areas adds interest to the results of our exercise.

The article is organized as follows. We next describe and summarize the existing empirical research on the effects of renewables on electricity prices, providing some general conclusions from this comprehensive overview. Section 3 deals with the price effects in the German-Austrian electricity market, with a description of the data, methods and results. Finally, Section 4 concludes and presents the main findings and implications.

2. Existing studies on renewable production and electricity prices

This section presents an overview of previous studies about the effect of renewables on electricity prices. As indicated above, any survey that attempts to draw general conclusions on this matter from disparate individual applications should be done with special care and should highlight both the differences and the global utility of each study. Therefore, we first describe the existing literature, grouping it in three categories: simulation-based approaches (Section 2.1), that is, studies based on simulation models that may use both real (past) or hypothetical data; empirical analyses (Section 2.2), based on real, past data, generally performed with econometric models; and finally other applications that provide some, but less concrete, information on the electricity price effects of renewables (Section 2.3). This is subsequently used to provide an overall picture and to draw some general messages from current knowledge in the field.

2.1. Simulation-based studies

There is now a rather sizeable literature that uses, for a number of countries and regions, simulation-based modeling approaches to deal with the effect of renewable production on electricity prices. We next enumerate and describe the most relevant pieces of research, which are grouped by country. Table 2 in Section 2.4 presents a summary of the main characteristics and findings of the papers mentioned in this section.
Starting with Germany, Sensfuss et al. (2008) use a model of the electricity market to run several simulations for situations with and without renewable production. They find that the average electricity price for Germany was reduced by 1.7 to 7.8 €/MWh due to the electricity production of renewables for the years 2001 and 2004 to 2006. In a subsequent application Sensfuss (2011) uses the same technique for the 2007-2010 period, showing that the 2010 effect is found to be at least (i.e. in a conservative calculation) between 5 and 6 €/MWh.

In another application, Bode and Groscurth (2006) construct a model of a synthetic electricity market by estimating demand and supply functions based on past experience and estimations of marginal costs of different production technologies (with basic assumptions and figures matching the 2005 actual situation). They quantify the effect at 0.5 to 0.6 €/MWh per GWh of additional renewable production (depending on assumptions about the demand function). Weber and Woll (2007) model the German electricity system for the following year, incorporating 34 technologies for electricity generation and prices of other energy goods and of CO₂ permits. Their calculations show that the hourly electricity prices are 4.04 €/MWh higher in a no-wind scenario with respect to the scenario with normal wind production. In a subsequent exercise, Weigt (2009) models the German electricity market to investigate the potential of wind generation to replace traditional fossil capacities. While doing so, he calculates electricity prices for the scenarios with and without wind generation, and finds lower prices for the former between January 2006 and June 2008. The price effect of wind generation also grows over time⁵. The study reports an average price effect of approximately -10 €/MWh for the studied period.

With a somewhat different approximation, Fürsch et al. (2012) calculate merit-order effects for Germany based on the DIME Model (Dispatch and Investment Model for Electricity Markets in Europe). This model is capable of reflecting international cross border flows and also allows for a changing electricity mix as a response to growing renewable participation in the mix. The results are forecasts for 2015, 2020, 2025 and 2030 that compare a scenario of the German energy market with a counterfactual of frozen renewable capacity at 2010 levels. As depicted in table 2,

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⁵ Off peak price effects: -2.4 €/MWh in 2006; -4.1 €/MWh in 2007; and -6.9 €/MWh in the first half of 2008. The price decreasing effect was found to be higher during peak hours: -10.1 €/MWh in 2006; -16.8 €/MWh in 2007; and -19.4 €/MWh in the first half of 2008.
the price differences between these two scenarios are found to be -2 €/MWh for 2015, -4 €/MWh for 2020, -5 €/MWh for 2025, and -10 €/MWh for 2030.\(^6\)

Finally, Traber and Kemfert (2009) develop an energy sector model based on an earlier work of Lise et al. (2006) where the European electricity sector incorporates the existing degree of competition and is modeled as a single market subject to the EU emissions trading scheme (ETS). In this setting they find that the German FIT system lowers electricity prices of German producers but also marginally increases consumer prices. This model is expanded in Traber and Kemfert (2011) to explicitly account for the consequences of the fluctuations in wind production. The paper finds that German market prices are roughly 3.7 €/MWh (or 5%) lower in their wind scenario than in the no-wind scenario. In another piece of research Traber et al. (2011) calculate the 2020 German wholesale electricity prices for a scenario with growing renewable shares and another one with frozen 2010 renewable capacity and expanded fossil capacity. Now the 2020 electricity wholesale price is expected to be 3.2 €/MWh lower for the scenario with growing renewables than it is for the frozen renewable scenario.

There are also some simulation studies on this question for Spain, another leading European supporter for renewables. Linares et al. (2008) use a simulation model of the electricity market to obtain results, up to 2020, for different scenarios: with and without a European carbon emission scheme and with or without additional national renewable support (or both). Given the actual existence of the EU ETS, that scenario may be used as counterfactual for the alternative situation where additional renewable support results in an expanded renewable capacity of 21.81 TWh in 2020. As a consequence, electricity prices would decrease by 1.74 €/MWh. In another simulation analysis, Sáenz de Miera et al. (2008) show that wind energy input significantly reduced Spanish electricity prices between 2005 and 2007 (amounts vary from -7.08 €/MWh to -12.44 €/MWh between the periods).

Lastly, Holttinen et al. (2001) carry out a simulation study for Nordpol (the Nordic electricity market that comprises Denmark, Finland, Norway and Sweden) to estimate the impact of wind generation on market prices. Using wind data from 1961 to 1990 to calibrate the model, in a 2010

\(^6\) The growth of the price effect, over time, is in line with the simultaneous growing difference in renewable generation between the considered scenarios.
forecast scenario the model yields spot price reductions of 2 €/MWh for each 10 TWh of additional wind production in the space of one week.

2.2. Empirical studies

In contrast to the preceding studies, several papers have taken advantage of the growing availability of ex-post data on electricity prices and renewable capacity in several countries to use different econometric approaches and techniques to compute the actual price effect of expanding renewables.

Starting again with Germany, although with very limited empirical evidence (that justifies carrying out our ad-hoc exercise later in the paper), Neubarth et al. (2006) set up a univariate regression model to investigate the effect of wind power production on day-ahead spot prices in Germany from September 2004 to August 2005. They find that the day-ahead electricity price falls by 1.89 €/MWh for each additional GW of wind production.

Using an empirical approach that will be largely replicated in our application with German data, Gelabert et al. (2011) use daily production quantities of different electricity generation types for Spain to investigate how they affect electricity prices during the 2005-2009 period. They find that each GW of additional renewable electricity production reduced Spanish electricity prices by roughly 2 €/MWh. A similar result was obtained by Sáenz de Miera et al. (2008), after picking three arbitrary days of February 2006 to perform an exhaustive comparison of electricity prices and wind energy production. They actually follow this approach to produce a ceteris paribus situation for all other influences except wind input, such as electricity demand, fuel prices, hydro production, etc. Finally, Gil et al. (2012) use a conditional probability approach to calculate how much the market penetration of wind generation influenced day-ahead electricity prices in Spain from 2007 to 2010. They indicate that electricity price reductions become increasingly likely when market penetrations of wind are higher. On average, they find that the electricity price without wind production is 9.7 €/MWh or 18% higher than it is with wind production.

Data availability and the interest to know the effects of an early rush to renewables, explain the various empirical applications to the Danish case. Ostergaard et al. (2006) analyze data of December 2005 on Danish electricity prices and wind generation. They distinguish between those
hours with wind and those hours without it. They find that Danish electricity prices would have been higher without any wind electricity generation (1 €/MWh in 2004; 4 €/MWh in 2005; and 2.5 €/MWh in 2006). A later application of Jonsson et al. (2010) uses data on day-ahead electricity prices and load and wind production forecasts for the Western Denmark market from January 2006 to October 2007. They find that wind production influenced prices heavily, detecting the strongest price-reducing effects during times of high wind production. Although they indicate that a quantification of the price effect is complicated, at some stage they assess the price differences between low wind (55-50 €/MWh during the day, 30 €/MWh at night) and high wind (30 €/MWh during the day, 18 €/MWh at night) situations. Such an approximate 40% electricity price variation between low-wind and high-wind circumstances is one of the highest in the literature. This is probably due to the unusually high renewable penetration and also the small size of the Danish system.

Nieuwenhout and Brand (2011) use wind and weather data from the Netherlands to reconstruct day-ahead wind generation figures for 2006-2009, and divide the data to create groups that correspond to low or no-wind production intervals. They find that average day-ahead prices at the Dutch electricity exchange were roughly 5% higher during the no-wind intervals with respect to the average of the entire sample for the analyzed period.

In another European case study, O'Mahoney and Denny (2011) employ 2009 hourly data for Ireland. They use a wide set of variables, including demand, wind generation and fossil fuel prices, to explain movements in the shadow price of electricity. As in other empirical applications they find that, during 2009, the price of electricity fell by 9.9 €/MWh per GW of wind.

As in some European countries, the effect of wind production on electricity prices has been heavily investigated in Texas due to the increasing relevance of renewables there. Two of the first exercises use high-frequency data (hourly and 15 min intervals, respectively), which in Nicholson et al. (2010) are used to analyze the 2007-2009 period with explanatory variables that include wind generation, production from gas plants, temperature, and past values of the electricity price. They find a range of decreasing effects of wind generation on balancing electricity prices of 0.67

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7 Ostergaard et al. (2006) is only available in Danish, but a summary of their findings can be obtained in Munksgaard and Morthorst (2008).
8 Jonsson et al. (2010) state that the Western Danish market has one of the highest wind penetrations in the world (occasionally exceeding 40%).
to 16.4 US$/MWh per additional GW of wind production (depending on the year, time of the day, and the area in the Texas network). Woo et al. (2011) study the 2007-2010 with a similar approximation, that includes nuclear generation, system load, price of gas, and a set of time dummies as additional explanatory variables, and finds that a 1 GWh increase in wind generation (during 15 minutes) decreased Texas balancing electricity prices between 13 and 44 US$/MWh. Actually, Zarnikau (2011) highlights that lacking transmission capacity within Texas has led to a situation where the intermittent wind production caused electricity prices to plunge in some regions while they peaked in others. Finally, Baldick (2011) argues that the high volatility of electricity prices in Texas also originates from the negative correlation between wind production and peak demand.

2.3. Studies with limited information on price effects

There are several other studies that, despite their interest in the electricity price effects associated to renewable production, do not yield a precise quantification. This usually happens because these papers typically focus on other questions and deal only marginally with the merit-order related effects of renewables on electricity prices. That is why we do not provide a detailed description of the papers, as in the previous sub-sections, but rather only a table that summarizes their main conclusions regarding this issue.

Table 1 thus presents the main conclusions of 17 papers that also differ from the previous literature in a much wider geographical coverage (including other US states, Australia or Israel). The table distinguishes, as before, between simulation-based and empirical approximations and shows the general price effect. Although it is now even harder to draw general conclusions from papers that are quite heterogeneous in methods, data and objectives, it seems quite obvious that electricity prices generally tend to fall due to increased renewable production. However, the magnitude of this effect varies greatly across studies, and there is disagreement about the duration and persistence of the effect.

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9 There is variability across the four existing zones in the Texas network, ranging from 1.3 to 4.4 €/MWh.
2.4. Comparability of existing studies

Sections 2.1 to 2.3, showed that higher renewable production was generally and consistently associated to reduced electricity prices, at least in the short term\(^\text{10}\). Yet the question is how much do electricity prices fall, that is, whether there is a common pattern in this phenomenon across the existing literature too. Table 2 provides a comprehensive summary of the different simulations and empirical approaches that have actually quantified the effects of renewable production on electricity prices, previously described in Sections 2.1 and 2.2.

\(^\text{10}\) Long-term effects may be different since the new generation mix and short-term lower prices influence investment decisions and thus future prices (Green and Vasilakos, 2010 and 2011; Milstein and Tishler, 2011; Gelabert et al., 2011).
Table 2. Literature on price effects of renewable generation

<table>
<thead>
<tr>
<th>(1) Paper</th>
<th>(2) Meth.</th>
<th>(3) Country</th>
<th>(4) Type</th>
<th>(5) Period</th>
<th>(6) Reported price change (€/MWh)</th>
<th>(7) As per</th>
<th>(8) During</th>
<th>(9) Common measure*</th>
<th>(10) Difference+</th>
</tr>
</thead>
<tbody>
<tr>
<td>Bode and Groscurth (2006)</td>
<td>S</td>
<td>All</td>
<td>Roughly 2005</td>
<td>2015-2016</td>
<td>1 GWh A G 1h</td>
<td>(-0.5, -0.6)</td>
<td>-</td>
<td>5.02</td>
<td></td>
</tr>
<tr>
<td>Fürsch et al. (2012)</td>
<td>S</td>
<td>Germany</td>
<td>All</td>
<td>2015-2030 (F)</td>
<td>A vs 2010 P</td>
<td>-0.40</td>
<td>-</td>
<td>11.42</td>
<td></td>
</tr>
<tr>
<td>Sensfuss (2011)</td>
<td>S</td>
<td>Germany</td>
<td>All</td>
<td>2007-2010</td>
<td>N vs N</td>
<td>-0.77</td>
<td>-</td>
<td>7.58</td>
<td></td>
</tr>
<tr>
<td>Neubarth et al. (2006)</td>
<td>E</td>
<td>Germany</td>
<td>Wind</td>
<td>2004-2005</td>
<td>1 GWh A G 1h</td>
<td>-1.89</td>
<td>-</td>
<td>16.32</td>
<td></td>
</tr>
<tr>
<td>Traber and Kemfert (2011)</td>
<td>S</td>
<td>Germany</td>
<td>Wind</td>
<td>2007-2008</td>
<td>N vs N</td>
<td>-0.80</td>
<td>-</td>
<td>4.65</td>
<td></td>
</tr>
<tr>
<td>Traber et al. (2011)</td>
<td>S</td>
<td>Germany</td>
<td>All</td>
<td>2020 (F) vs 2010</td>
<td>A vs 2010 P</td>
<td>-0.24</td>
<td>-</td>
<td>13.07</td>
<td></td>
</tr>
<tr>
<td>Weigt (2009)</td>
<td>S</td>
<td>Germany</td>
<td>Wind</td>
<td>2006-2008</td>
<td>N vs N</td>
<td>-1.78</td>
<td>-</td>
<td>3.51</td>
<td></td>
</tr>
<tr>
<td>Gelabert et al. (2011)</td>
<td>E</td>
<td>Spain</td>
<td>All</td>
<td>2005-2010</td>
<td>1 GWh A G 1h</td>
<td>-3.80</td>
<td>-</td>
<td>6.09</td>
<td></td>
</tr>
<tr>
<td>Linares et al. (2008)</td>
<td>S</td>
<td>Spain</td>
<td>All</td>
<td>2020 (F)</td>
<td>A vs A</td>
<td>-0.70</td>
<td>-</td>
<td>2.49</td>
<td></td>
</tr>
<tr>
<td>Holtinen et al. (2001)</td>
<td>S</td>
<td>Nordpool</td>
<td>Wind</td>
<td>2010 (F)</td>
<td>10 TWh A G 1 year</td>
<td>-2.0</td>
<td>-</td>
<td>2.60</td>
<td></td>
</tr>
<tr>
<td>Ostergaard et al. (2006)</td>
<td>E</td>
<td>Denmark</td>
<td>Wind</td>
<td>2004-2006</td>
<td>N vs N</td>
<td>-1.33</td>
<td>-</td>
<td>0.75</td>
<td></td>
</tr>
<tr>
<td>Nieuwenhout and Brand (2011)</td>
<td>E</td>
<td>Holland</td>
<td>Wind</td>
<td>2006-2009</td>
<td>N vs N</td>
<td>-6.17</td>
<td>-</td>
<td>0.43</td>
<td></td>
</tr>
<tr>
<td>O’Mahoney and Denny (2011)</td>
<td>E</td>
<td>Ireland</td>
<td>Wind</td>
<td>2009</td>
<td>1 MWh A G 1h</td>
<td>-9.90</td>
<td>-</td>
<td>-</td>
<td></td>
</tr>
<tr>
<td>Woo et al. (2011)</td>
<td>E</td>
<td>Texas</td>
<td>Wind</td>
<td>2007-2010</td>
<td>1 MWh A G 15 min</td>
<td>-2.34</td>
<td>-</td>
<td>-0.91</td>
<td></td>
</tr>
</tbody>
</table>

Notes: S, simulation; E, empirical approximation; F, forecast; A G, additional generation; A vs 2010 P, additional vs. 2010 production; N vs N, none vs. normal production; A vs A, additional vs. alternative production; L vs H, low vs. high production; * €/MWh for each additional GWh produced within 1 h; + difference in average hourly generation. Various sources have been used for the conversions: US Energy Information Administration, Red Eléctrica Española, Nordpoolspot, the German and Austrian electricity market and the reported papers. Source: The authors
Columns 6 to 8 in Table 2 report the price effects as provided by the particular studies of the surveyed literature. Some of the price changes refer to a certain increase in renewable or wind generation; others (mainly in the simulation studies) indicate the price difference between situations with normal wind and without wind. Obviously, the reported units are different and cannot be directly compared. We deal with this problem by recalculating the reported units so that all price changes are accounted in €/MWh and related to the increase in one GWh of renewable production (column 9 in Table 2). Finally, column 10 in Table 2 contains the average difference in generation per hour between scenarios so that the conversion process for studies that report price differences with respect to scenarios (not per amount of additional renewable generation) is transparent. Through this procedure, Table 2 yields comparable information that can be used for the conjoint evaluation of the different studies that will be developed in the following section. Yet, even with this additional effort, a comparative analysis and assessment of different studies should be carried out carefully and bear in mind the several remaining limitations.

Indeed, one factor that may render comparability difficult is precisely the homogenization effort because sometimes the additional data used to convert the results in comparable units may not coincide with the assumptions of the original studies. This is a likely outcome in simulation studies, where the overall generation and prices in the models might not coincide with actual values. Moreover, the analytical differences arise not only in terms of the use of real market data by empirical (vs. simulation) studies, but also within each of the general and apparently homogeneous approximations. For example, only some of the simulation studies use past weather data to estimate renewable output (e.g. Holttinen et al., 2001; Nieuwenhout and Brand, 2011), or consider the existence of an emission trading system (e.g. Sensfuss et al., 2008; Traber and Kemfert, 2011). Some of the simulation studies use a counterfactual with an amended power plant structure to represent possible long-run capacity mix adaptations (e.g. Traber et al., 2011; Sensfuss, 2011; Fürsch et al., 2012), while others leave the non-renewable capacity and generation values unchanged in their counterfactuals.

Significant heterogeneity also exists within empirical studies. For instance, a few papers use a high-frequency approach with an autoregressive element (e.g. Nicholson et al., 2010; Woo et al.,

11 The literature survey by Pöyry (2010) does such comparisons and can thus be misleading when evaluating merit-order and price effects.
Univariate approximations, that include only the renewable generation variable in the empirical analysis, are also common (e.g. Neubarth et al., 2006; Gil et al., 2012), and coexist with empirical exercises that include many other explanatory variables such as load, gas prices etc. (e.g. Nicholson et al. 2010; Woo et al, 2011; Gelabert et al., 2011). Given that additional variables may contain relevant explanatory power, it follows that univariate approaches are more likely to produce biased coefficients for the price effect of renewable generation and may prevent a meaningful comparative evaluation.

2.5. A comparative assessment of the literature

Even though the comparability of largely heterogeneous studies may be constrained by the reasons just mentioned, we believe that Table 2 provides as much comparable information as can be achieved in this area. In this section we use that information to gather some insights and general messages from existing studies on the effects of renewable production on electricity prices. In this sense, Figure 1 depicts column 9 of Table 2 in a graphical manner, grouping the results by countries or regions. The vertical axis on the right informs about the range of price effects, in €/MWh per additional GWh of renewable or wind energy produced, as reported for different countries by the specialized literature. The grey blocks indicate the number of studies per country or region.

Starting, as usual, with the evidence for Germany, existing studies seem to yield quite similar results albeit using different approaches and methods. Taken together, the nine studies on Germany report rather consistent results, quantifying the merit-order effect in a band between roughly -0.5 and -2 €/MWh, with most studies reporting reactions below -1 €/MWh. Only the studies by Neubarth et al. (2006) and Weigt (2009) are an exception that, at least for the former, may be explained by the use of a univariate approach that may produce biased estimates (see Section 2.4). The first results reported for Germany are for 2001 (Sensfuss et al., 2008), with forecasts available until 2030 (Fürsch et al. 2012), and do not indicate a growing price effect when measured in €/MWh for each GWh of additional renewable production12.

12 When reported as a price difference between full and no renewable generation, however, the price difference grows over time. This is logical, due to the increased capacity and participation of renewables in the energy system.
The four Spanish studies yield a price effect of approximately -2 €/MWh, without significant variation even though the methodologies of Sáenz de Miera et al. (2008), Gelabert et al. (2011) and Gil et al. (2012) differ substantially. However, the trends over the years reported by Sáenz de Miera et al. (2008) and Gelabert et al. (2011) are rather different: whereas the former detect a strong increase in the merit-order effect between 2006 and 2007, the latter show a significant decrease during the same period. In this sense, at least compared to the literature on Germany, the studies on Spain seem to give somewhat less coherent results.

The highest price effects were actually detected in small electricity markets like Denmark, the Netherlands, and Ireland. This is explained by the accounting unit for price effects, €/MWh for each additional GWh of renewable energy produced, as one GWh of production is a much smaller part of large electricity markets. For example, the Danish and Dutch hourly wind production in 2006 (a year analyzed by Nieuwenhout and Brand (2011), Ostergaard et al. (2006) and Jonsson et al. (2010)) was far below 1 GW on average; for Ireland it was even less. On the contrary, Spanish and German hourly wind 2006 production was 2.6 GW and 3.5 GW
respectively, about three to five times bigger than it was in their smaller counterparts. This means that, accounting for market sizes, price effects are not that different in Europe\textsuperscript{13}.

Results for the Scandinavian Nordpool market (Holttinen et al., 2001) are similar in size as the results for Germany and Spain (the price effect of 1,7 €/MWh is in between the average Spanish and German effect). Total size of the Nordpool market is also in between these two markets. When relating the results of Holttinen et al. (2001) for the Nordpool market to those of other markets one must consider the methodological differences of this particular study.\textsuperscript{14} One might have expected lower price effects of wind power in this particular Nordpool market because of the availability of reservoir hydro capacity that can adjust to the volatility of renewables without raising the costs of the marginal power plant in the system. However, due to the unique methodology in Holttinenen et al. (2001) and the availability of only one study for the joint Nordpool market, we are reluctant to speculate on possible reasons of the outcome on the Nordpool market.

Finally, the Texas studies are the least comparable because of acute methodological differences (high-frequency autoregressive models by Nicholson et al. (2010) and Woo et al. (2011)), and the strong asymmetries between ERCOT, Texas electricity system, and its European counterparts. The variation of the price effect in the Texas studies is high among different zones within ERCOT, which in some cases lack interconnection capacity (Zarnikau, 2011). Baldick (2011) also states that the typical wind patterns in Texas are particularly unfavorable and this makes even more complicated to compare the Texas results with other international applications.

Summarizing then, the comparison of results per country/region shows that the particular estimates of merit-order effects do not show as many differences among countries as the individual results suggest at first sight (columns 6-8, Table 2). When converted into homogeneous units, at least the European merit-order effects are quite comparable in size.

Given the joint analysis of previous results, it is particularly interesting to see how these results relate to the biggest European electricity market, which also features an important scheme for renewable support. Therefore, we next present new results for Germany and Austria, an integrated electricity market in the core of Europe that also strongly influences neighboring areas.

\textsuperscript{13} This is also illustrated in columns 6 and 7 of Table 2 for the cases where the quantification of the price effect is expressed as a difference between normal and zero renewable generation (which is not available for all studies).

\textsuperscript{14} Holttinen et al. (2001) is the only study that uses a weekly frequency of data, calibrates wind output in the model in a unique way, and allows for capacity changes between scenarios in areas outside Nordpool which are interconnected.
Our new empirical analysis also closes a research gap since, as observed in previous subsections, the effect of renewable production on electricity prices in Germany and Austria were mainly analyzed through simulation approaches in the past. The following section presents novel results in an extensive data framework with various model variations that provide new insights for a number of relevant merit-order effect questions, including for instance those related to the German nuclear abandonment program.

3. An empirical analysis for Austria and Germany

This section presents a new empirical exercise on the effects of renewable production on the price of electricity based on real data for Germany and Austria, both of which are analyzed together given the high degree of integration between their electricity markets. We first show some basic facts regarding the German and Austrian electricity market (Section 3.1), followed by a description of the data and empirical methodology used in the study (Section 2.2). The final sub-section deals with the results from the estimation of a multivariate regression model, which are later compared to the outcomes of existing applications.

In the preceding parts of the paper we observed that existing empirical literature on Germany regarding this matter is limited, so any contribution in this area may prove useful to complete and understand that literature. Moreover, we feel that this paper closes a significant research gap for several reasons. First, the German and Austrian market possesses many of the characteristics that make merit-order effect analyses especially interesting, that is, it has a large share of renewables and a strong renewable support scheme. Second, the economic importance of this area coupled with its central location in Europe (with 13 bordering countries) make the contagion of the price effects associated to renewable production on neighboring countries likely. Third, its high integration with adjacent economies and electricity markets, add further relevance to the exercise. Finally, the results can be extrapolated to other big economies that might opt for a similar path with high renewable shares as predicted by future energy scenarios (see e.g. IPCC, 2011; IEA, 2012).
3.1. The German and Austrian electricity market

The German part of the market is much greater in size (roughly eight to one) and defines the combined market as the largest in Europe. Taken together, Germany and Austria feature a significant renewable production capacity, although most solar and wind capacity is installed in Germany (in early 2012 29,700 wind MW in Germany vs. 1,150 MW in Austria and 24,500 solar MW in Germany vs. 107 MW in Austria). This difference is related to a much higher carbonization rate in the German electricity generation system (mostly coal and lignite) while Austria is endowed with significant hydro capacities. In this setting, climate change and other environmental concerns explain the early introduction (1991) of a FIT system for renewable sources, with the exception of hydro, in Germany. The support system gave grid access to all renewable producers but its success in triggering significant investments in renewable capacities was limited. This led to the 2000 reform through the Renewable Energy Act (Erneuerbare Energien Gesetz, EEG), which guaranteed fixed tariffs to different renewable technologies and plant sizes (that now includes small hydro). The Act has continuously been amended ever since, with a prominent and controversial reduction in solar remuneration since early 2012, although changes have not affected the basic mechanics of the 2000 EEG. Table A.1 in the Appendix reproduces the financial efforts related to the German FIT and the EEG that, in plain contrast with Austria\textsuperscript{15}, converted Germany into a world leader in photovoltaic and wind installed capacity (Bode and Groscurth, 2006).

Figures 2 (Germany) and 3 (Austria) summarize the recent evolution of the electricity mix in both countries, with the right-hand side illustrating the evolution of renewables. Carbon-intensive technologies clearly prevail in Germany, even though the system participation of renewables has grown significantly in the last few years (renewable production tripled from 40 TWh per year in 2001 to 120 TWh in 2011). The diagrams clearly illustrate the importance of renewables in the German-Austrian electricity system, where they provide a much higher contribution to electricity generation than in other global markets.

\textsuperscript{15} In 2003 the Austrian government launched its “ecological electricity regulation” which has frequently been amended ever since. Yet this legislation only sets goals for the installed capacity of hydro, wind, biomass, and waste for 2020, and it does not include extensive market interventions such as those introduced by Germany’s EEG.
3.2. Data and methods

We construct a multivariate regression model similar to Gelabert et al. (2011). The electricity price ($P_{elec}$) is the dependent variable, and the explanatory variables are the demand of electricity ($Load$), the renewable electricity production from solar and wind ($RE$), the gas price ($P_{gas}$), and the exports and imports of electricity ($Exlm$), i.e.,

$$\Delta P_{elec,t} = \beta_0 + \beta_1 \Delta Load_t + \beta_2 \Delta RE_t + \beta_3 \Delta P_{gas,t} + \beta_4 \Delta Exlm_t + \beta_5 \times dummies + \epsilon_t$$ (1)
where $\Delta$ represents the first difference operator, subscript $t$ stands for time (daily sequence), and $\epsilon$ is the standard error term. Additionally, the model includes several dummies to control for typical time patterns in electricity markets: daily dummies capture the usual fluctuations during the week (e.g. the differences between a working day and a weekend); monthly dummies capture seasonal patterns and a third dummy variable is one in supra-regional bank holidays and zero otherwise.

We use the day-ahead price for electricity (Phelix) at the Electricity Exchange in Leipzig (EEX), where electricity for Germany and Austria is traded. The day-ahead price was chosen over the intra-day price because it represents a larger share of trading volume. Obviously, explanatory variables must be assembled to match the characteristics of the day-ahead electricity price. Ideally, they are also day-ahead forecasts because the foreseen values of load and production data are the basis for purchasing and selling decisions on the day-ahead market\textsuperscript{16}. Consequently, the data on renewable production forecasts consist of solar and wind production forecasts as reported by several sources: EEX and the five grid operators (Tennet, EnBW Transnet, 50Hertz and Amprion for Germany; and APG for Austria). Hence, the renewable production aggregate includes German solar production (given the negligible Austrian solar capacity), and German and Austrian wind production.

Several studies found system load/demand variation to be highly relevant for price formation\textsuperscript{17}, which is included as a second explanatory variable in this analysis. We use the forecasted load in the German-Austrian power system, that is, the sum of all deliveries from the transmission system through directly connected transformers and lines to distribution networks and end consumers. This load variable would typically include non-renewable large-scale production, while significant parts of wind and solar production would never enter the transmission grid in the first place. This variable therefore combines nicely with the renewable production variable (see above) and minimizes any possible double-counting biases between the load and the solar-wind variables. Data was assembled via the European Network of Transmission System Operators for Electricity (ENTSOE) and APG.

\textsuperscript{16} Gomez-Quiles and Gil (2011) also follow this procedure with a similar line of reasoning.
\textsuperscript{17} Ostergaard et al. (2006); Sensfuss et al. (2008); Weigt et al. (2009); Jonsson et al. (2010); Gelabert et al. (2011); Woo et al. (2011).
Typically, the price of other energy goods is another valuable price indicator for electricity. Austria and Germany have highly developed economies where energy markets are linked either through substitution possibilities for consumers or through input factor influences (such as gas-fired power plants). Thus, the price of natural gas was included as an explanatory variable to represent such cross-influences\textsuperscript{18}. The day-ahead gas price data originates from the two German gas transport companies: Gaspool and NetConnect Germany.

Moreover, Germany and Austria both have several international electricity transmission nodes with 10 of their 13 neighbors\textsuperscript{19} (without considering the nodes that directly connect Germany and Austria). As international electricity trade can mitigate national demand and supply fluctuations, transmission capacity must be taken into account when estimating price effects\textsuperscript{20}. Therefore, all electricity flows from/to the German-Austrian area to/from the ten neighboring countries with transmission capacity were considered and a transmission aggregate data series was composed. This electricity export and import data originates from the European Market Coupling Company GmbH and the five electricity grid German and Austrian operators. Unfortunately, the international electricity flow data was not available as forecasts\textsuperscript{21}.

Due to data availability the period of investigation is from 1 July 2010 to 30 June 2012\textsuperscript{22}. Figure A.1 in the Appendix plots the data used in the empirical analysis. The augmented Dickey-Fuller (ADF) test was applied to test for unit roots and stationarity. Most of the series are I(1) so first differences of the data series are taken for computing the model. The ADF test results are represented in the Appendix, along with a correlation matrix. This indicates a negligible probability of problems related to multicollinearity in the estimations\textsuperscript{23}. Daily averages were taken in all cases.

\textsuperscript{18} Alternatively the price of coal or lignite could also be included, but past research on Germany and Austria indicates that gas prices are more relevant for electricity prices (see Redl and Haas, 2007; Obersteiner and Redl, 2007; Sensfuss et al., 2008).

\textsuperscript{19} Denmark, Sweden, Poland, Czech Republic, Hungary, Slovenia, Italy, Switzerland, France, and the Netherlands.

\textsuperscript{20} Fürsch et al. (2012) argue that international transmission is relevant for merit-order effects, whereas Traber and Kemfert (2009) find that the German EEG did influence electricity prices in EU neighboring countries.

\textsuperscript{21} Forecasts regarding international electricity trade are only available for the capacity of transmission nodes. However, the variable of interest is not the available capacity of transmission nodes but the real flow. Therefore, the electricity flow data was used here even though it consists of realized values instead of day-ahead forecasts.

\textsuperscript{22} German law requires data on production of wind and solar plants to be published since July 2010. Prior reliable data is not available, thus limiting the data period to two years.

\textsuperscript{23} See for example Verbeek (2008, p. 43) who gives the rule of thumb that multicollinearity is very unlikely to be a problem if the correlations between the explanatory variables do not exceed 0.8 (or -0.8 respectively). The only value close to this boundary is between load and the electricity price, the latter being the dependent variable. Therefore, this value does not reflect possible multicollinearity issues among explanatory variables.
to eliminate unwanted ad-hoc anomalies and short-term noise. This means that prices (gas and electricity) are the daily averages of the day in question, whereas quantity data (renewable production, load, and exports-imports) are the hourly average of the 24 hours for each day.

3.3. Results

The final model of Equation 1 is constructed in various steps. First we estimate a model with the dummies as the only explanatory variables (Model 1), showing how much seasonal and weekday patterns matter for electricity prices. The electricity quantity variable Load is added for the estimation of Model 2. Model 3 is the complete regression with all the variables of Equation 1, although results are presented for several variations (Models 3A to 3F) in a procedure that provides further details on specific issues and also serves as a robustness test. The first variation (Model 3A and 3B) is a distinction between the first year of the data period (1 July 2010 to 30 June 2011) and the second year (1 July 2011 to 30 June 2012) to observe possible differences due to varying penetrations of renewable sources (Figures 1 and 2 indicate strong growth in renewable production since 2010) and to possible long-run adjustment of the electricity sector to earlier merit-order effects (see Sections 1 and 2). Moreover, the separation of the first and last year roughly coincides with the German nuclear exit, where half of the German nuclear capacity was switched off immediately and the so-called energy transition (Energiewende) was announced. Therefore, this model also provides information on whether such changes in capacity mix affected the price effects of renewable production.

A second model variation (Models 3C and 3D) implements a reduction of the dataset to the upper quarter of high-load days and the lower quarter of low-load days. This is done to verify the hypothesis that renewable production has a much higher impact on electricity prices when the electricity system is closer to full capacity as observed by other papers (e.g. Ostergaard et al., 2006; Jonsson et al., 2010; Gelabert et al., 2011). The third variation involves a separation of solar and wind electricity sources through the use of separate coefficients that intend to identify the different production patterns of these technologies. The last variation (Model 3F) estimates

24 Also, Neubarth et al. (2006) found that, when daily values are taken instead of 15 or 60-minute intervals, forecasts of renewable production (wind) are more relevant for the definition of day-ahead prices at EEX.
Equation 1 using weekly average instead of daily averages. The results for all model variations are reported in Table 3.

Table 3. OLS estimation of daily changes in electricity prices

<table>
<thead>
<tr>
<th>Model 1</th>
<th>Model 2</th>
<th>Model 3</th>
<th>Model 3A</th>
<th>Model 3B</th>
<th>Model 3C</th>
<th>Model 3D</th>
<th>Model 3E</th>
<th>Model 3F</th>
</tr>
</thead>
<tbody>
<tr>
<td>( \Delta P_{\text{elec}, t} )</td>
<td>( \Delta P_{\text{elec}, t} )</td>
<td>( \Delta P_{\text{elec}, t} )</td>
<td>1st year</td>
<td>2nd year</td>
<td>High load</td>
<td>Low load</td>
<td>RE split up</td>
<td>Weekly</td>
</tr>
<tr>
<td>( \Delta \text{Load}_t )</td>
<td>0.000727</td>
<td>0.000243</td>
<td>0.000514</td>
<td>0.000154</td>
<td>0.000345</td>
<td>0.000538</td>
<td>0.000242</td>
<td>0.00059</td>
</tr>
<tr>
<td>(0.000117)</td>
<td>(0.000083)</td>
<td>(0.000164)</td>
<td>(0.000084)</td>
<td>(0.000148)</td>
<td>(0.000193)</td>
<td>(0.000084)</td>
<td>(0.000172)</td>
<td></td>
</tr>
<tr>
<td>( \Delta \text{RE}_t )</td>
<td>-0.00103</td>
<td>-0.000963</td>
<td>-0.001036</td>
<td>-0.00109</td>
<td>-0.00093</td>
<td>-0.000853</td>
<td></td>
<td></td>
</tr>
<tr>
<td>(0.000081)</td>
<td>(0.000133)</td>
<td>(0.000112)</td>
<td>(0.000084)</td>
<td>(0.000148)</td>
<td>(0.000193)</td>
<td>(0.000084)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>( \Delta \text{Wind}_t )</td>
<td>-0.00103</td>
<td>-0.000963</td>
<td>-0.001036</td>
<td>-0.00109</td>
<td>-0.00093</td>
<td>-0.000853</td>
<td></td>
<td></td>
</tr>
<tr>
<td>(0.000081)</td>
<td>(0.000133)</td>
<td>(0.000112)</td>
<td>(0.000084)</td>
<td>(0.000148)</td>
<td>(0.000193)</td>
<td>(0.000084)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Adjusted R²</td>
<td>0.51</td>
<td>0.591</td>
<td>0.72</td>
<td>0.726</td>
<td>0.737</td>
<td>0.746</td>
<td>0.54</td>
<td>0.72</td>
</tr>
<tr>
<td>D-W</td>
<td>2.429</td>
<td>2.562</td>
<td>2.557</td>
<td>2.557</td>
<td>2.566</td>
<td>2.082</td>
<td>2.444</td>
<td>2.564</td>
</tr>
<tr>
<td>Observations</td>
<td>731</td>
<td>731</td>
<td>731</td>
<td>365</td>
<td>366</td>
<td>183</td>
<td>183</td>
<td>731</td>
</tr>
<tr>
<td>Daily dummies</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
<td>no</td>
</tr>
<tr>
<td>Monthly dummies</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
</tr>
<tr>
<td>Holiday dummy</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
<td>yes</td>
<td>no</td>
</tr>
</tbody>
</table>
| Notes: All models include an intercept. Bold print indicates a coefficient significantly different from zero (p<0.05). The adjusted R-squared assumes acceptable values throughout Model 3. D-W critical values for all versions of Models 3 (except 3F) are respectively: DL 1.51 to 1.72; DU 2.28 to 2.49. As a consequence, standard errors (in parenthesis below the coefficients) are robust to heteroskedasticity and serial correlation (Newey and West, 1987). Source: The authors

In line with the theory of the merit-order effect and previous research (see Sections 1 and 2), the coefficient of the renewable production variable is always negative and significant. Ceteris paribus, day-ahead electricity prices for Germany and Austria decrease by roughly 1 €/MWh for each additional expected GW produced by renewable sources (solar and wind). The price effect of renewable production is stable at around -1 €/MWh throughout all versions of model 3: for high or low load days, for the first year or the last year of the analyzed period, and for individual or

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25 Actually, the p-value of the RE variable in model 3 is lowest of all four explanatory variables at 0.000.
conjoint representation of solar and wind technologies. This price decreasing effect of renewables is also within the range of earlier (mostly simulation) studies of Germany, which range approximately in between -0.5 and -2.5 €/MWh. The merit-order effect found in our analysis roughly corresponds to an electricity price decrease of 2%, since the average price was about 48 €/MWh for the period under study.

The different variations of Model 3 reveal other valuable information. When comparing the first and last years in 3A and 3B, the extent of the influence of the renewable production variable is quite similar in both periods, with a very low statistical probability of the renewable coefficients being different in the two periods which contrasts with the findings of other applications (see Table 2). This result is also interesting in light of the nuclear exit in mid 2011, when seven German nuclear plants were switched off after the Fukushima event, as the data reveal no immediate effect of that change in the electricity mix on the size of the merit-order effect. Other variables provide further information on this matter: while the load appears to be decisive for prices (as expected with a positive coefficient) during the first year, in the last year the gas price assumes more informative value for electricity prices with a shift in the significance of coefficients. Hence, the fact that fossil capacities had to substitute the deactivated nuclear capacity is reflected in the significance of gas prices, even though it did not translate into a change in the size of the merit-order effect. Yet another parallel development does not show in the results of Model 3A and 3B: German electricity imports also rose as a consequence of the nuclear exit, but the coefficient of the export/import variable remains insignificant. Future research in this area may find different results for the changes over time when new data becomes available.

More interesting results arise from the comparison of high and low-load days (Model 3C and 3D). The effect of renewable production on the electricity price appears to be more pronounced for high-load days. This result is in line with previous evidence (e.g. Ostergaard et al., 2009; Senfuss et al., 2008, Weigt et al., 2009; Woo et al., 2011), although it is not very robust in our case because standard deviations do not allow for the conclusion that the coefficient of renewable production of high and low-load days is different. One possible explanation for this phenomenon could be a relatively constant slope in the German electricity supply curve in the high-quantity

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26 The increased electricity imports (particularly from the Czech Republic, France and Nordpool) are depicted in Figure A.1 in the Appendix, with the export/import aggregate falling in April 2011 and slowly recovering until the end of 2011. The coefficient of this aggregate comes closest to significance for model 3B (also in the expected way, with a negative sign).
area. Another reason for this divergence may be found in the other explanatory variables: the gas price is only significant for high-load days because of additional requirements for fossil fuels so that peak-load plants can cope with the unusually high demand. At the same time, this shift in significances of the gas price variable already captures differences between high and low-load days, which then do not materialize in the coefficients of renewable production anymore. As a matter of fact, the coefficients of the renewable production variable for low and high-load days show bigger differences when the gas price variable is omitted in Model 3C and 3D27.

Model 3E yields no significant differences between the coefficients for solar and wind production and therefore the price effects associated to the wind or solar source of renewable production seem to be very similar28. It is also worthwhile to note that the other coefficients of the remaining explanatory variables in model 3E are almost identical to those in the base model, which did not distinguish between solar and wind production. Finally, the results of Model 3F indicate that the importance and magnitude of renewable production remain unchanged when the frequency of the data is altered. The renewable coefficient is somewhat smaller, which is no big surprise because valuable information may be lost when expanding the intervals to weekly observations.

In sum, the estimation results are quite robust across models. Significant coefficients always have the same sign and the effects of renewable production on prices are also consistent: the day-ahead electricity price for the German-Austrian market falls by around 1 €/MWh for each additional GWh of expected renewable electricity production. Our findings are quite compatible with earlier research on merit-order effects for Germany, which reported results in the 0.5-2.3 €/MWh interval (see Section 2). Given that the only existing empirical study for Germany, by Neubarth et al. (2006), is somewhat outdated and applies a univariate approach that does not contemplate the effects of capacity use, exports and imports or gas prices, we feel that this paper contributes to and complements the existing literature on the matter29.

27 The coefficients are -0.001018 (high load) and -0.000834 (low load) and show the largest discrepancy throughout all model variations. Apparently the gas price and load variables overlap to some extent, although they do so without causing problematic multicollinearity (see the Appendix).
28 Note that the effects could be quite different with a higher data frequency (i.e., distinguishing between different times of the day). In this case, the effect of solar production is potentially higher because its production pattern coincides with demand peaks. Yet, such intra-day specialties cannot be captured or represented in a study based on daily averages.
29 See e.g. Green and Vasilakos (2010) and Twomey and Neuhoff (2010).
Yet we would like to emphasize, once again, that the results of this and other studies on the electricity price effects of renewables should be interpreted carefully. Some factors, such as the market structure of the power sector, were not considered in this exercise and may be potentially relevant\textsuperscript{30}. Adding more detailed data about the remaining (non-renewable) mix in electricity generation may also affect the reported estimations. Finally, data availability is problematic (especially for renewable production data prior to July 2010) and thus severely constrains the period of analysis. Future research using data beyond 2012 could overcome this problem and provide valuable information on the temporal evolution of the effect of renewable production on German and Austrian electricity prices.

4. Conclusions

Supporting renewables to achieve a less polluting and (foreign) dependent energy sector has many consequences. Among them, an increased renewable production of electricity crowds out other high(ger) marginal-cost technologies and results in lower electricity prices. Quantifying this theoretical merit-order effect is not simple, however, and multiple analytical techniques have been used for this purpose. Applied research in this area has mainly been carried out through simulation and empirical approximations that have generally reported seemingly large discrepancies regarding the size of this effect.

This paper starts by providing a comprehensive review of the increasing literature in this area, so that common patterns and trends can be identified. In this sense, when results are converted into homogeneous units (€/MWh per each additional GWh of renewable production), the smallest merit-order effects are found in large European markets (Nordpool \textsuperscript{-0.03}; Germany \textsuperscript{-0.24 to -2.83}; Spain \textsuperscript{-1.1 to -3.99}), in contrast to much higher price effects in small markets (Netherlands \textsuperscript{-6.17}; Denmark \textsuperscript{-1.33 to -9.87}; Ireland \textsuperscript{-9.9}). This is not surprising, as 1 GWh of additional production is a larger share of small electricity markets. When accounting for differences in market size, our comparison of existing European studies shows an interesting and novel pattern as merit-order effects are now quite similar. This is the case for those electricity systems with sizeable fossil capacities (Germany, Spain, Holland, Ireland, and Denmark), where fossil plants are still the

\textsuperscript{30} Moreover, our study for Germany considers the virtually complete integration with Austria, including Austrian data for all explanatory variables. Given that the electricity exchange in Leipzig treats Germany and Austria as one market with one electricity price, Austrian production data should not be ignored.
price-setting marginal plant, at least during demand peaks. Even the Nordpool market with its high reservoir hydro capacities shows price effects similar to markets of comparable sizes as Germany or Spain. The differences and/or similarities of merit order effects through different countries leave room for future more detailed research, for example about the pattern of peak hours in markets, or fuel import costs.

The second part of the paper presents new results on the price effects of renewable production in the German-Austrian market. A major reason for this exercise is the limited empirical research on this matter, which is clearly anomalous for an area with the biggest electricity market in Europe and with high renewable shares and intense renewable support schemes. One of the major findings of this research is that the day-ahead electricity price fell by roughly 1 €/MWh (around 2% of the electricity price) for each GWh of average hourly predicted renewable electricity generation between July 2010 and June 2012. The effect was found to be stable throughout different model variations, and the reported figure is in line with previous research results with simulation approximations.

Other interesting conclusions arise from our German-Austrian empirical exercise. First, the changed electricity mix brought about by the German nuclear exit did not affect the size of the merit-order effect, which remained stable at 1 €/MWh before and after the deactivation of seven nuclear plants by mid-2011. Second, the price effect of wind and solar was more or less equal, which is rather counterintuitive because solar is expected to have a higher impact for its coinciding production with demand peaks. This may be due to the chosen frequency of the analysis, which does not reflect alternating production patterns within a day. Another relevant result refers to the weak evidence on stronger merit-order effects at times of high electricity demand. This may, however, be due to the fact that fossil fuel use in peak-load plants (which is another indicator of the relevance of demand peaks) does affect electricity prices.
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References


APPENDIX

Table A.1. Costs of the German EEG renewable support

<table>
<thead>
<tr>
<th></th>
<th>2004</th>
<th>2005</th>
<th>2006</th>
<th>2007</th>
<th>2008</th>
<th>2009</th>
<th>2010</th>
</tr>
</thead>
<tbody>
<tr>
<td>Total EEG generation (GWh)</td>
<td>24,969</td>
<td>28,417</td>
<td>38,511</td>
<td>43,967</td>
<td>51,545</td>
<td>67,010</td>
<td>73,348</td>
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<td>Total cost of EEG (€ Mio)</td>
<td>3,611</td>
<td>4,498</td>
<td>5,810</td>
<td>7,879</td>
<td>9,016</td>
<td>10,774</td>
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<td>By generation type</td>
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<tr>
<td>% of EEG Generation</td>
<td>12</td>
<td>11.3</td>
<td>9.6</td>
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<td>7</td>
<td>6.5</td>
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<tr>
<td>Costs (€ Mio)</td>
<td>337.67</td>
<td>364.10</td>
<td>366.56</td>
<td>417.70</td>
<td>378.81</td>
<td>382.38</td>
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<td>Costs (€/kWh)</td>
<td>7.307</td>
<td>7.329</td>
<td>7.408</td>
<td>7.510</td>
<td>7.378</td>
<td>7.838</td>
<td>8.335</td>
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<td>% of EEG Generation</td>
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<td>7.1</td>
<td>5.4</td>
<td>4.1</td>
<td>3.1</td>
<td>2.7</td>
<td>1.44</td>
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<tr>
<td>Costs (€ Mio)</td>
<td>182.17</td>
<td>219.24</td>
<td>195.62</td>
<td>192.88</td>
<td>155.87</td>
<td>142.64</td>
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<td>Costs (€/kWh)</td>
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<td>7.023</td>
<td>7.028</td>
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<tr>
<td>% of EEG Generation</td>
<td>13.6</td>
<td>16.8</td>
<td>21.1</td>
<td>23.8</td>
<td>26.6</td>
<td>30.6</td>
<td>31.16</td>
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<td>Costs (€ Mio)</td>
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<td>1,337.37</td>
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<td>% of EEG Generation</td>
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<td>0.03</td>
<td>0.05</td>
<td>0.06</td>
<td>2.64</td>
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<tr>
<td>% of EEG Generation</td>
<td>66.3</td>
<td>61.9</td>
<td>59.6</td>
<td>59.3</td>
<td>57</td>
<td>51.4</td>
<td>46.42</td>
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<tr>
<td>Costs (€ Mio)</td>
<td>2,300.48</td>
<td>2,440.68</td>
<td>2,733.77</td>
<td>3,508.44</td>
<td>3,561.04</td>
<td>3,388.90</td>
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<tr>
<td>Costs (€/kWh)</td>
<td>9.010</td>
<td>8.968</td>
<td>8.899</td>
<td>8.829</td>
<td>8.518</td>
<td>8.785</td>
<td>8.850</td>
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</tr>
<tr>
<td>% of EEG Generation</td>
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<td>n.a.</td>
<td>n.a.</td>
<td>n.a.</td>
<td>n.a.</td>
<td>0.22</td>
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<td>n.a.</td>
<td>n.a.</td>
<td>n.a.</td>
<td>26.06</td>
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<td>Costs (€/kWh)</td>
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<td>n.a.</td>
<td>n.a.</td>
<td>n.a.</td>
<td>n.a.</td>
<td>14.679</td>
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<td></td>
<td></td>
</tr>
<tr>
<td>% of EEG Generation</td>
<td>1.4</td>
<td>2.9</td>
<td>4.3</td>
<td>4.6</td>
<td>6.2</td>
<td>8.8</td>
<td>14.48</td>
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<tr>
<td>Costs (€ Mio)</td>
<td>282.65</td>
<td>679.11</td>
<td>1,176.8</td>
<td>1,597.48</td>
<td>2,218.62</td>
<td>3,156.52</td>
<td>5,089.94</td>
</tr>
<tr>
<td>Costs (€/kWh)</td>
<td>52.424</td>
<td>53.262</td>
<td>53.094</td>
<td>51.825</td>
<td>48.787</td>
<td>47.792</td>
<td>43.559</td>
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</tbody>
</table>

Notes: The German EEG sets tariffs through a complex system depending on energy type and specific plant characteristics (size, year of construction etc), which are also adjusted frequently. Rates of €/kWh are the averages within each category. Geothermal tariffs have been between 14.5 and 24 €/kWh since 2009.
Source: German Grid Operators (Informations Plattform der Deutschen Übertragungsnetzbetreiber, EEG KWK-G).
Some seasonal patterns are visible in the series of Figure A.1, along with a high volatility in renewable production (RE) and some scarcity issues during some exceptionally cold days in the beginning of 2012, which are visible in the electricity and gas price series (Phelix Day and PGas).
As typical for time series data, there are non-stationary patterns in the datasets. We tested for unit roots in the data by use of the augmented Dickey-Fuller test (Dickey and Fuller, 1979). Several versions of the test have been run with altering conditions for included lags and the inclusion of a trend. They indicate that some series are I(1). Hence, first differences of the data series are taken for the model calculus. The ADF test results are depicted in Table A.2.

### Table A.2. Augmented Dickey Fuller test statistics (probabilities)

<table>
<thead>
<tr>
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<th>Levels</th>
<th>1st differences</th>
</tr>
</thead>
<tbody>
<tr>
<td>Phelix Day-ahead Electricity Price (Pelec)</td>
<td>0.0982</td>
<td>0.000</td>
</tr>
<tr>
<td>Renewable Production Forecast (RE)</td>
<td>0.0000</td>
<td>0.000</td>
</tr>
<tr>
<td>Load forecast (Load)</td>
<td>0.1057</td>
<td>0.000</td>
</tr>
<tr>
<td>Export Import Aggregate (Exlm)</td>
<td>0.3298</td>
<td>0.000</td>
</tr>
<tr>
<td>Natural Gas Price (Pgas)</td>
<td>0.0001</td>
<td>0.000</td>
</tr>
</tbody>
</table>

Note: ADF test results (p-values), where lag length is determined by the Akaike Info Criterion and a trend variable is included.
Source: The authors

Table A.3 reports the correlation matrix for the daily changes (first differences) of the data series. The matrix shows no perfect (or excessive) collinearities, indicating that the risk of multicollinearity in the estimations is low.

### Table A.3. Correlations of daily changes

<table>
<thead>
<tr>
<th></th>
<th>EXIM</th>
<th>LOAD</th>
<th>PGAS</th>
<th>PHELIX_DAY</th>
<th>RE</th>
</tr>
</thead>
<tbody>
<tr>
<td>EXIM</td>
<td>1</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>LOAD</td>
<td>-0.3760</td>
<td>1</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>PGAS</td>
<td>-0.0230</td>
<td>0.1812</td>
<td>1</td>
<td></td>
<td></td>
</tr>
<tr>
<td>PHELIX_DAY</td>
<td>-0.4720</td>
<td>0.7481</td>
<td>0.2175</td>
<td>1</td>
<td></td>
</tr>
<tr>
<td>RE</td>
<td>0.6020</td>
<td>-0.2669</td>
<td>-0.0093</td>
<td>-0.4450</td>
<td>1</td>
</tr>
</tbody>
</table>

Note: The correlations reported correspond to the variables in first differences.
Source: The authors

---

31 See for example Verbeek (2008, p. 43) who gives the rule of thumb that multicollinearity is very unlikely to be a problem if the correlations between the explanatory variables do not exceed 0.8 (or -0.8 respectively). The only value close to this boundary is between load and the electricity price, the latter being the dependent variable.